Lecture 8 - Reasoning under Uncertainty (Part II)

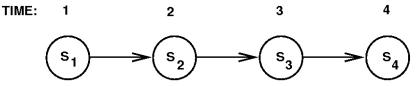
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June 21, 2022

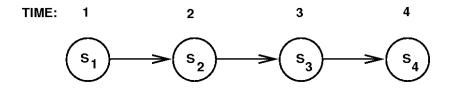
Readings: Poole & Mackworth (2nd ed.)Chapt. 8.5 - 8.9

Probability and Time

- A node repeats over time
- explicit encoding of time
- chain has length = amount of time you want to model
- event-driven times or clock-driven times
- e.g. Markov chain



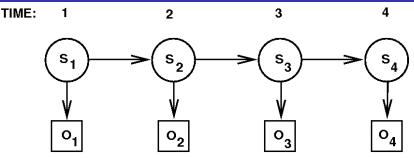
Markov assumption



$P(S_{t+1}|S_1,\ldots,S_t)=P(S_{t+1}|S_t)$

This distribution gives the dynamics of the Markov chain

Hidden Markov Models (HMMs)



Add: observations O_t (always observed, so the node is square) and

observation function $P(O_t|S_t)$

Given a sequency of observations O_1, \ldots, O_t , can estimate filtering:

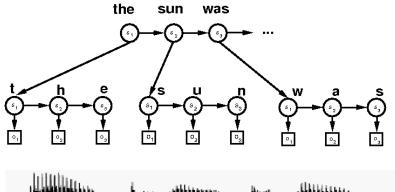
 $P(S_t|O_1,\ldots,O_t)$

or smoothing, for k < t

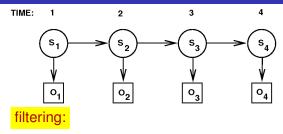
 $P(S_k|O_1,\ldots,O_t)$

Speech Recognition

- Most well known application of HMMs
- observations : audio features
- states : phonemes
- dynamics : models e.g. co-articulation
- HMMs : words
- Can build hierarchical models (e.g. sentences)

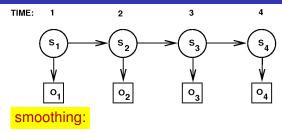


Belief Monitoring in HMMs



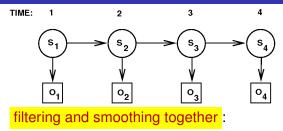
$$\begin{aligned} \alpha_{i} &= P(S_{i}|o_{0}...,o_{i}) \\ &\propto P(S_{i},o_{0},...,o_{i}) \\ &= P(o_{i}|S_{i}) \sum_{S_{i-1}} P(S_{i},S_{i-1},o_{0},...,o_{i-1}) \\ &= P(o_{i}|S_{i}) \sum_{S_{i-1}} P(S_{i}|S_{i-1}) P(S_{i-1},o_{0},...,o_{i-1}) \\ &\propto P(o_{i}|S_{i}) \sum_{S_{i-1}} P(S_{i}|S_{i-1}) \alpha_{i-1} \end{aligned}$$

Belief Monitoring in HMMs



$$\begin{aligned} \beta_{i+1} &= P(o_{i+1} \dots, o_T | S_i) \\ &= \sum_{S_{i+1}} P(S_{i+1}, o_{i+1}, \dots, o_T | S_i) \\ &= \sum_{S_{i+1}} P(o_{i+1} | S_{i+1}, o_{i+2}, \dots, o_T, S_i) P(S_{i+1}, o_{i+2}, \dots, o_T | S_i) \\ &= \sum_{S_{i+1}} P(o_{i+1} | S_{i+1}) P(o_{i+2}, \dots, o_T | S_{i+1}, S_i) P(S_{i+1} | S_i) \\ &= \sum_{S_{i+1}} P(o_{i+1} | S_{i+1}) P(S_{i+1} | S_i) \beta_{i+2} \end{aligned}$$

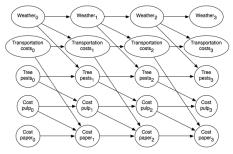
Belief Monitoring in HMMs



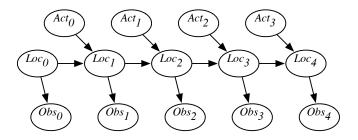
$$\alpha_i \beta_{i+1} = P(o_{i+1} \dots, o_T | S_i) P(S_i | o_0 \dots, o_i) \propto P(S_i | O)$$

Dynamic Bayesian Networks (DBNs)

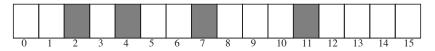
- in general, any Bayesian network can repeat over time: DBN
- Many examples can be solved with variable elimination,
- may become too complex with enough variables
- event-driven times or clock-driven times



- Suppose a robot wants to determine its location based on its actions and its sensor readings: Localization
- This can be represented by the augmented HMM:



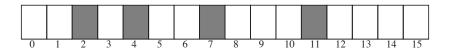
• Circular corridor, with 16 locations:



- Doors at positions: 2, 4, 7, 11.
- Noisy Sensors
- Stochastic Dynamics
- Robot starts at an unknown location and must determine where it is, known as the kidnapped robot problem.
- see handout robotloc.pdf

- *P*(*Observe Door* | *At Door*) = 0.8
- P(Observe Door | Not At Door) = 0.1

- $P(Loc_{t+1} = I | Action_t = goRight \land Loc_t = I) = 0.1$
- $P(Loc_{t+1} = l + 1 | Action_t = goRight \land Loc_t = l) = 0.8$
- $P(Loc_{t+1} = I + 2 | Action_t = goRight \land Loc_t = I) = 0.074$
- P(Loc_{t+1} = l'|Action_t = goRight ∧ Loc_t = l) = 0.002 for any other location l'.
 - All location arithmetic is modulo 16.
 - The action goLeft works the same but to the left.



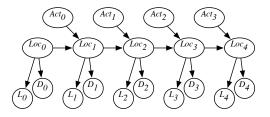
observe door, go right, observe no door, go right, observe door where is the robot?

 $P(Loc_2 = 4 | O_0 = d, A_0 = r, O_1 = \neg d, A_1 = r, O_2 = d) = 0.42$

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Combining sensor information

- Example: we can combine information from a light sensor and the door sensor Sensor Fusion
- Key Point: Bayesian probability ensures that evidence is integrated proportionally to its precision.
- Sensors are precision weighted



 Loc_t robot location at time t D_t door sensor value at time t L_t light sensor value at time t

Probability Distribution and Monte Carlo



John von Neumann 1903 - 1957



Stanlislaw Ulam 1909-1984

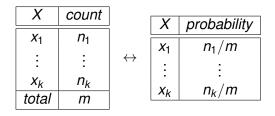


ENIAC 1949



Monte Carlo 1949

- Idea: probabilities \leftrightarrow samples
- Get probabilities from samples:

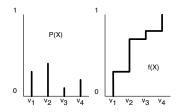


 If we could <u>sample</u> from a variable's (posterior) probability, we could <u>estimate</u> its (posterior) probability.

Generating samples from a distribution

For a variable X with a discrete domain or a (one-dimensional) real domain:

- Totally order the values of the domain of *X*.
- Generate the cumulative probability distribution : $f(x) = P(X \le x)$.
- Select a value y uniformly in the range [0, 1].
- Select the *x* such that f(x) = y.



p is true probability, *s* is sample average, *n* is number of samples

•
$$P(|s-p| > \epsilon) \leq 2e^{-2n\epsilon^2}$$

if we want an error greater than
 ϵ in less than a fraction δ of the cases, solve for *n*:

$$2e^{-2n\epsilon^2} < \delta$$
 $n > rac{-lnrac{\delta}{2}}{2\epsilon^2}$

• we have

$\epsilon \; \mathrm{error}$	cases with error $> \epsilon$	samples needed
0.1	1/20	184
0.01	1/20	18,445
0.1	1/100	265

Forward sampling in a belief network

- Sample the variables one at a time ;
- sample parents of X before you sample X.
- Given values for the parents of *X*, sample from the probability of *X* given its parents.
- for samples s_i , $i = 1 \dots N$:

$$P(X = x_i) \propto \sum_{s_i} \delta(x_i) = N_{X = x_i}$$

where

$$\delta(x_i) = \begin{cases} 1 & \text{if } X = x_i \text{ in } s_i \\ 0 & \text{otherwise} \end{cases}$$

Sampling for a belief network: inference

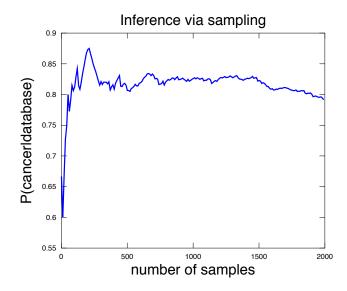
Sample	Malfnction	Cancer	TestB	TestA	Report	Database
<i>S</i> ₁	false	false	true	true	false	false
S 2	false	true	true	true	true	true
S 3	false	true	true	true	true	true
<i>S</i> ₄	false	false	false	true	false	false
S 5	true	true	true	true	false	false
<i>s</i> ₆	false	true	false	true	false	false
S 7	false	false	false	true	false	true
<i>s</i> ₁₀₀₀	false	false	false	true	false	false

To get $P(H = h_i | E = e_i)$ simply

- count the number of samples that have $H = h_i$ and $E = e_i$, $N(h_i, e_i)$
- divide by the number of samples that have $E = e_i$, $N(e_i)$
- $P(H = h_i | E = e_i) = \frac{P(H = h_i \land E = e_i)}{P(E = e_i)} = \frac{N(h_i, e_i)}{N(e_i)}$
- P(C = True | Database = True) based on first 7 samples?

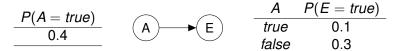
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Forward Sampling



- To estimate a posterior probability given evidence $Y_1 = v_1 \land \ldots \land Y_j = v_j$:
- If, for any *i*, a sample assigns Y_i to any value other than v_i reject that sample.
- The non-rejected samples are distributed according to the posterior probability.
- in the Hoeffding bound, n is the number of non-rejected samples

Example Network



If we draw *N* samples $s_{i=1...N}$ by

- sampling A: a_{i=1...N}
- sampling from E given A: e_{i=1...N}

then

- $\approx N_t = 0.4N$ of them will have A = true, and of these $\approx 10\%$ will have E = true
- $\approx N_f = 0.6N$ of them will have A = false, and of these $\approx 30\%$ will have E = true

Example Network

$$\begin{array}{c} \underline{P(A = true)} \\ \hline 0.4 \end{array} \qquad (A \longrightarrow E) \qquad \begin{array}{c} A & P(E = true) \\ \hline true & 0.1 \\ false & 0.3 \end{array}$$

so we have

$$\frac{A \quad E \quad N_{AE}}{\text{true false } N_{tf} = 0.4 \times 0.9 \times N}$$

$$\text{true true } N_{tt} = 0.4 \times 0.1 \times N$$

$$\text{false false } N_{ff} = 0.6 \times 0.7 \times N$$

$$\text{false true } N_{ft} = 0.6 \times 0.3 \times N$$
We want to compute

$$P(a|e) = P(A = true|E = true) \propto \sum_{s_i} \delta(a_i = true) \delta(e_i = true)$$

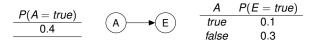
$$P(a|e) = \frac{P(a \land e)}{P(e)} \approx \frac{N_{tt}}{N_{tt} + N_{ft}}$$
$$= \frac{0.1 \times 0.4 \times N}{0.1 \times 0.4 \times N + 0.3 \times 0.6 \times N} = 0.182$$

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Α	P(E = true)
true	0.1
false	0.3

- we can do better since we can weight the samples
- weights = prob. that the evidence is observed
- N_t samples with A = true have weight of w_t = 0.1 this is P(E = true|A = true)
- N_f samples with A = false have weight of $w_f = 0.3$ this is P(E = true | A = false)
- can do better because we don't need to generate the 90% of samples (when A = true) that don't agree with the evidence we simply assign all samples a weight of 0.1
- thus, we are mixing exact inference (the 0.1) with sampling.



• Compute sum of all weights of the samples with *A* = *true*

$$W_t = \sum_i w_t \delta(a_i = true) = N_t \times 0.1$$

• Compute sum of all weights of the samples with *A* = *false*

$$W_f = \sum_i w_f \delta(a_i = false) = N_f \times 0.3$$

finally, compute

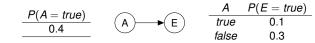
$$P(a|e) = \frac{W_t}{W_t + W_f} = \frac{0.1 \times 0.4 \times N}{0.1 \times 0.4 \times N + 0.3 \times 0.6 \times N}$$



- In fact, the As don't need to even be sampled from P(A)
- Can be sampled from some q(A), say q(A = true) = 0.5
- and each sample will have a new weight P(a)/q(a)
- q(A) is a proposal distribution.
- helps when it is hard to sample from P(A), but we can evaluate P^{*}(A) ∝ P(A) given a sample (see slide 24)



- $\begin{array}{c|c} A & P(E = true) \\ \hline true & 0.1 \\ false & 0.3 \end{array}$
- rejection sampling uses q = P
- rejection sampling uses all variables including observed ones, and all weights on samples are set to 1.0
- $N'_t = q(a)N$ samples with A = true have weight of $0.1 \times \frac{P^*(a)}{q(a)} = 0.1 \times \frac{\alpha 0.4}{0.5}$
- $N'_f = q(\overline{a})N$ samples with A = false have weight of $0.3 \times \frac{P^*(\overline{a})}{q(\overline{a})} = 0.3 \times \frac{\alpha 0.6}{0.5}$



• total weight of all samples with *A* = *true*

$$W'_t = \sum_i w_i \delta(a_i = true) = N'_t \times 0.1 \times \frac{\alpha 0.4}{0.5}$$
$$= 0.5N \times 0.1 \times \frac{\alpha 0.4}{0.5} = 0.1 \times \alpha \times 0.4 \times N$$

• total weight of all samples with *A* = *false*

$$egin{aligned} \mathcal{W}_f' &= \sum_i w_i \delta(a_i = \textit{true}) = \mathcal{N}_f' imes 0.3 imes rac{lpha 0.6}{0.5} \ &= 0.5 \mathcal{N} imes 0.3 imes rac{lpha 0.6}{0.5} = 0.3 imes lpha imes 0.6 imes \mathcal{N} \end{aligned}$$

$$\frac{P(A = true)}{0.4} \qquad (A) \rightarrow (E) \qquad \frac{A \quad P(E = true)}{true \quad 0.1}$$

$$false \qquad 0.3$$

• finally, compute

$$P(a|e) = \frac{W'_t}{W'_t + W'_f} = \frac{0.1 \times \alpha \times 0.4 \times N}{0.1 \times \alpha \times 0.4 \times N + 0.3 \times \alpha \times 0.6 \times N}$$

When drawing samples gets hard

- Why is it hard to sample from P(A)?
- because you need to know $P^*(A) \propto P(A)$ for all values of A (to normalize properly)
- in high dimensional spaces, there can be a lot
- consider the task of measuring the average (or maximum) depth of this lake - how do you draw samples? You cannot miss the canyons!



Figure 29.3. A slice through a lake that includes some canyons.

•
$$P(A) = \frac{P^*(A)}{\sum_A P^*(A)}$$

•
$$\sum_{A}$$
 is potentially intractable

- consider if A is continuous or discrete-valued over 500 dimensions.
- causes problems for exact inference as well

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- Sometimes we may want to choose a proposal distribution that is different than the actual probability distribution
- We may want to skew the proposal because we may have some additional knowlege about the data, for example
- or, we can generate proposals from the data itself using some procedural knowledge that is not directly encoded in the BN
- Can be important in multiple/many dimensions,

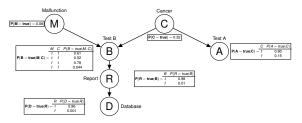
Recall variable elimination: To compute $P(Z, Y_1 = v_1, ..., Y_j = v_j)$, we sum out the other variables, $Z_1, ..., Z_k = \{X_1, ..., X_n\} - \{Z\} - \{Y_1, ..., Y_j\}$. $P(Z, Y_1 = v_1, ..., Y_j = v_j)$

$$=\sum_{Z_k}\cdots\sum_{Z_1}\prod_{i=1}^n P(X_i|parents(X_i))_{Y_1=V_1,\ldots,Y_j=V_j}$$

Now, we sample Z_{l+1}, \ldots, Z_k and sum Z_1, \ldots, Z_l ,

$$= \sum_{s_i = \{z_{l+1,i},...,z_{k,i}\}} \left[\sum_{Z_1...Z_l} \prod_{i=1}^l P(Z_i | parents(Z_i))_{Y_1 = v_1,...,Y_j = v_j} \right] \frac{P(Z_{l+1,i},...,Z_{k,i})}{q(z_{l+1,i},...,z_{k,i})}$$

Importance Sampling example



Compute P(B|D = true, A = false) by sampling *C* and *M*.

• use
$$q(C = true) = 0.5$$

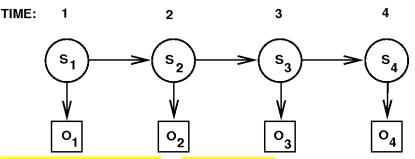
and $q(M = true) = P(M = true) = 0.08$

• use
$$q(C = true) = q(M = true) = 0.5$$

$$P(B|D = true, A = false) \propto \sum_{s_i \in \{c_i, m_i\}} P(B, D = true, A = false|c_i, m_i)$$

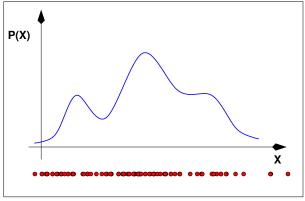
see sampling-inference.pdf

Stochastic Sampling for HMMs (and other DBNS)

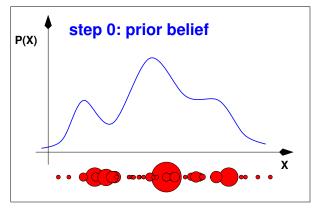


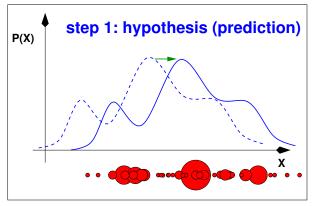
Sequential Monte Carlo or Particle Filter

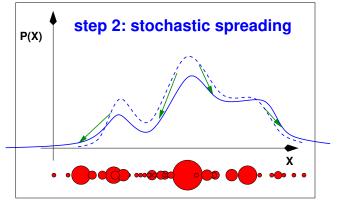
- sequential stochastic sampling
- keep track of P(S_t) at the current time t
- represent $P(S_t)$ with a set of samples
- update as new observations o_{t+1} arrive
 - 1. predict $P(S_{t+1}) \propto P(S_{t+1}|S_t)$
 - 2. compute weights as $P(o_{t+1}|S_{t+1})$
 - 3. resample according to weights

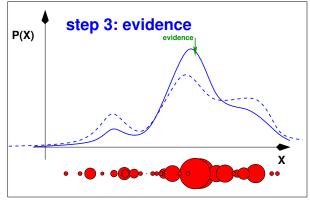


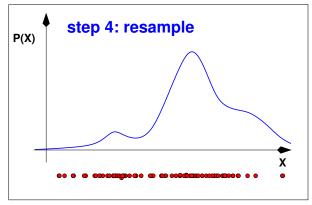
sample i: {x, }



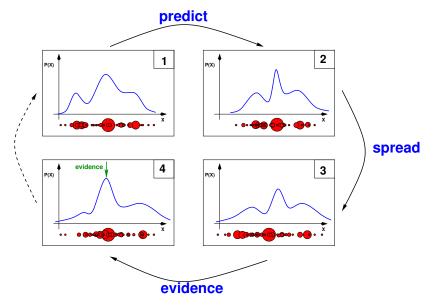




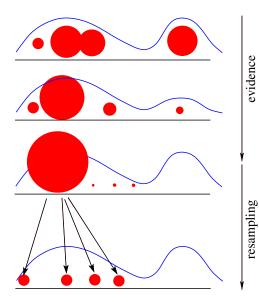




Bayesian Sequential Updates



Resampling



 avoids
 degeneracies in the samples

- all importance weights \rightarrow 0 except one
- performance of the algorithm depends on the

resampling method

 Supervised Learning under Uncertainty (Poole & Mackworth (2nd ed.)10.1,10.4)